

# ■ Εκτιμητική

## □ Σημειακές Εκτιμήτριες και οι κατανομές τους

### ➤ για τη μέση τιμή

- ✓ παρατηρήσεις από κανονική κατανομή με γνωστή διασπορά
- ✓ «πολλές» παρατηρήσεις από οιαδήποτε κατανομή

### ➤ για την διασπορά

- ✓ παρατηρήσεις από κανονική κατανομή με άγνωστη μέση τιμή

### ➤ για τη μέση τιμή

- ✓ παρατηρήσεις από κανονική κατανομή με άγνωστη διασπορά

- Έστω  $X_1, X_2, \dots, X_n$  ανεξάρτητες και ισόνομες τ.μ. με  $E(X_i) = \mu$  και  $\Delta(X_i) = \sigma^2$   $i=1, \dots, n$ .

Τότε ορίσω των τ.μ.  $\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$   
 συν. τον μέσο όρο τους

$$E(\bar{X}) = \frac{1}{n} E(X_1 + X_2 + \dots + X_n) = \frac{1}{n} [E(X_1) + \dots + E(X_n)]$$

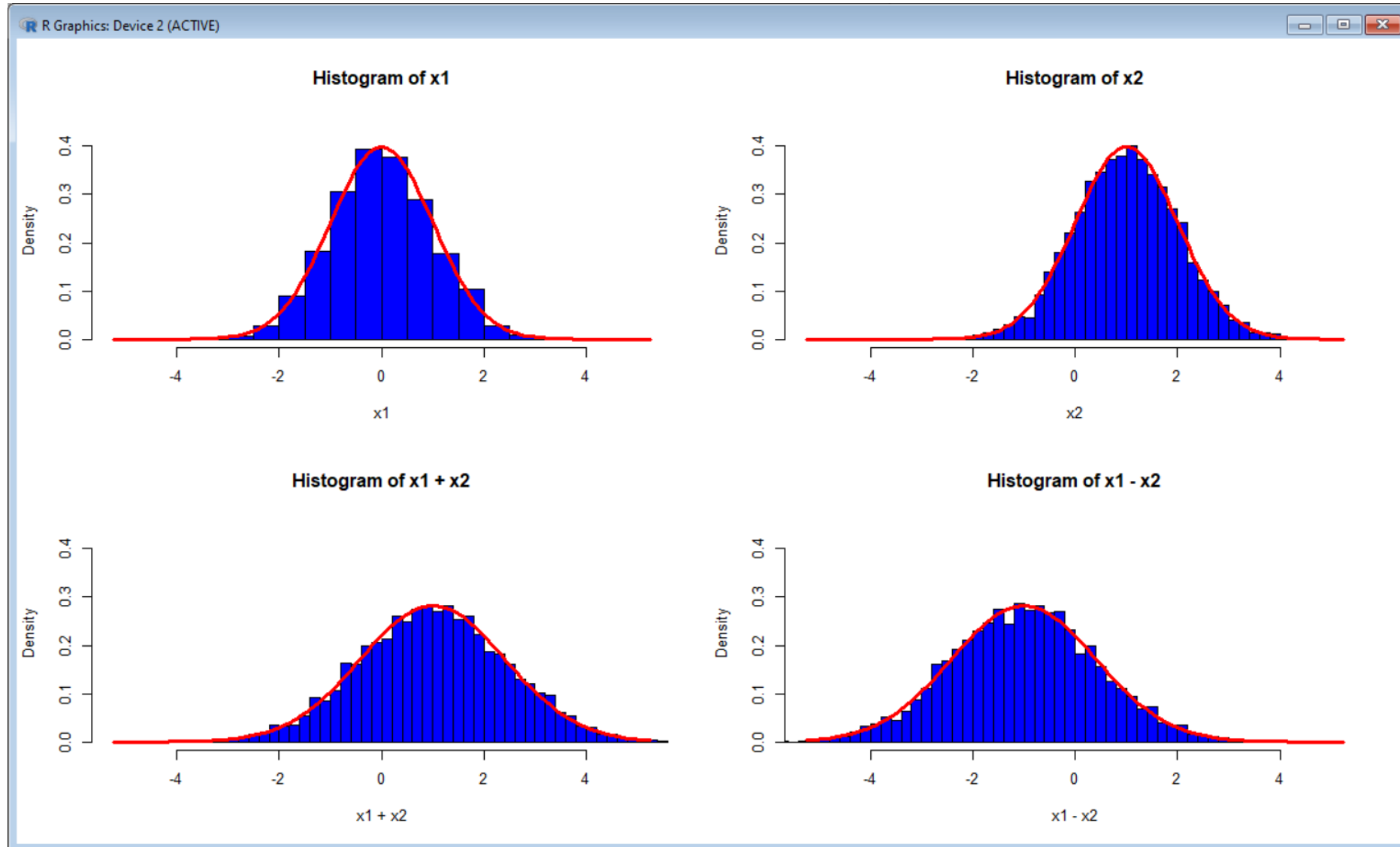
$$= \frac{1}{n} \cdot n \cdot \mu = \mu$$

$$\Delta(\bar{X}) = \frac{1}{n^2} \Delta(X_1 + X_2 + \dots + X_n) \stackrel{\text{ανεξ.}}{=} \frac{1}{n^2} (\Delta(X_1) + \dots + \Delta(X_n))$$

$$= \frac{1}{n^2} \cdot n \sigma^2 = \frac{\sigma^2}{n}$$

• εάν οι τ.μ.  $X_1 \sim N(\mu_1, \sigma_1^2)$  και  $X_2 \sim N(\mu_2, \sigma_2^2)$  είναι ανεξάρτητες τότε για την τ.μ.  $Y = X_1 \pm X_2$  ισχύει  $Y \sim N(\mu_1 \pm \mu_2, \sigma_1^2 + \sigma_2^2)$

# Γραμμικοί συνδυασμοί κανονικών τυχαίων μεταβλητών



## Γραμμικοί συνδυασμοί κανονικών τυχαίων μεταβλητών

```
> x1<-rnorm(5000)
> x2<-rnorm(5000,1,1)
> par(mfrow=c(2,2))
> c(mean(x1),var(x1),sd(x1))
[1] 0.002814884 0.997127619 0.998562777
> c(0,1,1)
[1] 0 1 1
> hist(x1,nclass=25,col=4,prob=TRUE,xlim=c(-5.25,5.25),ylim=c(0,0.45))
Waiting to confirm page change...
> curve(dnorm(x),lwd=3,col="red",add=TRUE)
>
> c(mean(x2),var(x2),sd(x2))
[1] 1.004343 1.017108 1.008518
> c(1,1,1)
[1] 1 1 1
> hist(x2,nclass=25,col=4,prob=TRUE,xlim=c(-5.25,5.25),ylim=c(0,0.45))
> curve(dnorm(x,1,1),lwd=3,col="red",add=TRUE)
>
> c(mean(x1+x2),var(x1+x2),sd(x1+x2))
[1] 1.007158 2.019812 1.421201
> c(1,2,sqrt(2))
[1] 1.000000 2.000000 1.414214
> hist(x1+x2,nclass=45,col=4,prob=TRUE,xlim=c(-5.25,5.25),ylim=c(0,0.45))
> curve(dnorm(x,1,sqrt(2)),lwd=3,col="red",add=TRUE)
>
> c(mean(x1-x2),var(x1-x2),sd(x1-x2))
[1] -1.001528 2.008659 1.417272
> c(-1,2,sqrt(2))
[1] -1.000000 2.000000 1.414214
> hist(x1-x2,nclass=45,col=4,prob=TRUE,xlim=c(-5.25,5.25),ylim=c(0,0.45))
> curve(dnorm(x,-1,sqrt(2)),lwd=3,col="red",add=TRUE)
> par(mfrow=c(1,1))
```

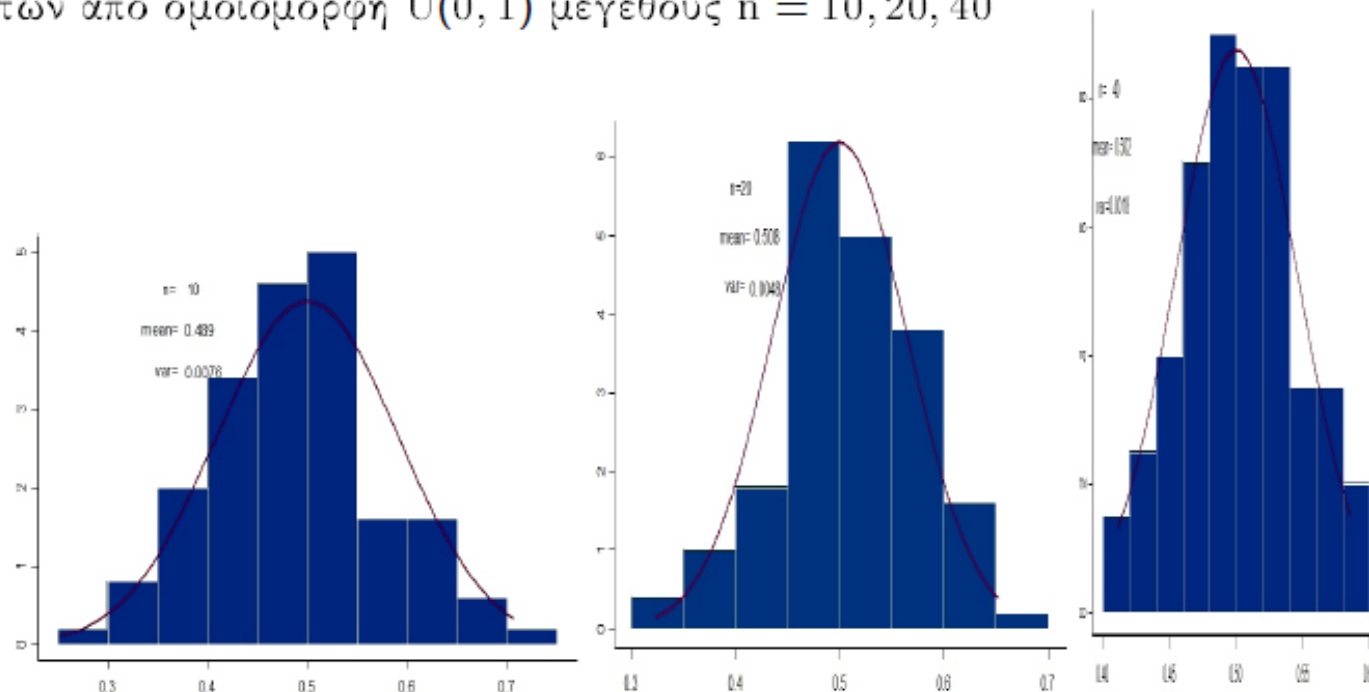
## Κεντρικό Οριακό Θεώρημα

- Έστω  $X_1, X_2, \dots, X_n$  ανεξάρτητες και ισόνομες τ.μ. με  $E(X_1) = \mu < \infty$  και  $0 < \text{Var}(X_1) = \sigma^2 < \infty$  τότε για τον δειγματικό μέσο ισχύει

$$\overline{X}_n \xrightarrow[n \rightarrow \infty]{\text{κατά κατανομή}} N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{ή} \quad \frac{\overline{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \xrightarrow[n \rightarrow \infty]{\text{κ. κ.}} N(0, 1)$$

δηλαδή  $\lim_{n \rightarrow \infty} P\left(\frac{\sqrt{n}(\overline{X}_n - \mu)}{\sigma} \leq x\right) = \Phi(x) \quad \forall x \in \mathbb{R}$

- ιστογράμματα σχετικών συχνοτήτων (100 παρατηρήσεων) των μέσων τιμών τυχαίων δειγμάτων από ομοιόμορφη  $U(0, 1)$  μεγέθους  $n = 10, 20, 40$



**Άσκηση** Οι μετρήσεις του ουρικού οξέως (σε mg/100ml) για άνδρες ηλικίας 35-50 ετών ακολουθούν κανονική κατανομή με μέση τιμή 5.4 και τυπική απόκλιση 1.

- α. Να βρεθεί η πιθανότητα για έναν άνδρα αυτής της ηλικίας η μέτρηση του ουρικού οξέως να είναι μεγαλύτερη από 5, όταν γνωρίζουμε ότι αυτή είναι μικρότερη του μέσου.
- β. Παίρνουμε δείγμα 25 μετρήσεων. Ποια είναι η πιθανότητα ο μέσος όρος των μετρήσεων να είναι μεταξύ 5.3 και 5.6;

**Παρατήρηση:** Για το ερώτημα β. δεν είναι αναγκαία η υπόθεση ότι οι μετρήσεις ακολουθούν (προέρχονται) από κανονική κατανομή. Καθώς το πλήθος του,  $n=25$ , είναι «οριακά» μεγάλο (αν και θα επιτυγχάναμε καλύτερη προσέγγιση με περισσότερες παρατηρήσεις) μπορεί με ασφάλεια να χρησιμοποιηθεί το κεντρικό οριακό θεώρημα και το ερώτημα επιλύεται όποια και να είναι η κατανομή των  $X_i$

β. Παίρνουμε δείγμα 25 μετρήσεων. Ποια είναι η πιθανότητα ο μέσος όρος των μετρήσεων να είναι μεταξύ 5.3 και 5.6;

†  
εάν  $X_1, \dots, X_{25}$  ανεξάρτητες παρατηρήσεις από  $N(\mu, \sigma^2)$ .

τότε  $\bar{X} = \frac{X_1 + \dots + X_{25}}{25} \sim N\left(\mu, \frac{\sigma^2}{25}\right)$  δηλαδή  $N\left(5.4, \frac{1}{25}\right) \equiv N\left(5.4, \frac{1}{5^2}\right)$

και τυποποιούμε

$$\begin{aligned} P(5.3 < \bar{X} < 5.6) &= P\left(\frac{5.3 - 5.4}{\frac{1}{5}} < \frac{\bar{X} - 5.4}{\frac{1}{5}} < \frac{5.6 - 5.4}{\frac{1}{5}}\right) \\ &= P(-0.5 < Z < 1) = \Phi(1) - \Phi(-0.5) \\ &= \underbrace{\Phi(1)}_{\substack{N(0,1) \text{ συμμ. ως προς } 0 \\ \Phi(1) = 1 - \Phi(-1)}} - (1 - \Phi(0.5)) \\ &= 0.841 - 1 + 0.691 = 0.532 \end{aligned}$$

Κατανομή των στατιστικών συνδυασμών  $\bar{X}$  και  $S^2$

$$1) X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$\bar{X}$  ως γραμμικός συνδυασμός κανονικών γ.μ. αυτοθουθεί κανονική κατανομή

$$E(\bar{X}) = E(X_i) = \mu \quad \text{var}(\bar{X}) = \frac{\text{var}(X_i)}{n} = \frac{\sigma^2}{n}$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\hat{\mu} = \bar{X} \quad \left. \begin{array}{l} \text{αμερόληπτος εκτιμητής του } \mu \\ \text{και } \lim_{n \rightarrow \infty} \text{var}(\bar{X}) = \lim_{n \rightarrow \infty} \frac{\sigma^2}{n} = 0 \end{array} \right\} \begin{array}{l} \text{συνεής} \\ \text{εκτιμητής} \\ \text{του } \mu. \end{array}$$

- Εάν οι παρατηρήσεις είναι πολλές ( $n \gg 25$ ) από όποια κατανομή και εάν προέρχονται οι  $X_i$ , ισχύει το Κεντρικό Οριακό Θεώρημα και  $\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$ .



# Ποσότητες οδηγοί - ποσοστιαία σημεία των αντίστοιχων κατανομών

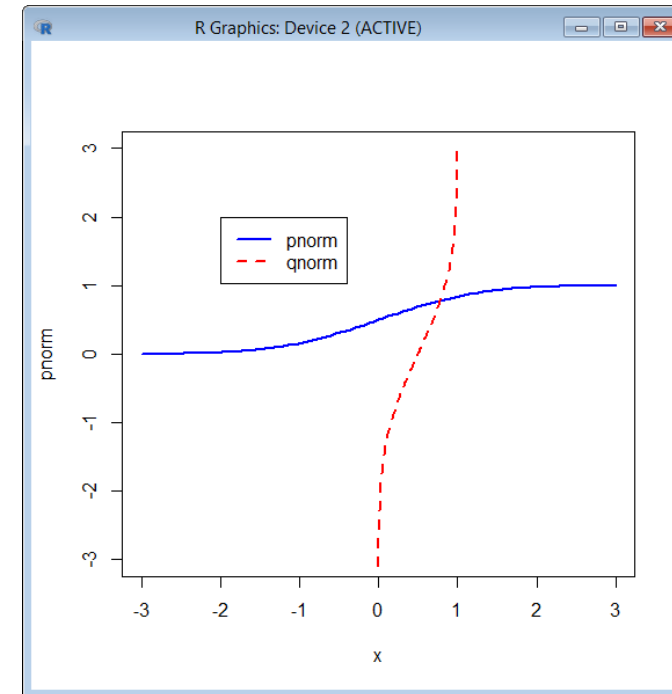
$X_1, X_2, \dots, X_n$  τυχαίο δείγμα από  $N(\mu, \sigma^2)$

$$\bar{X} = \frac{\sum_{i=1}^n X_i}{n} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \Rightarrow \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1)$$

Η α.σ.κ. της  $N(0,1)$ ,  $\Phi(x)$ , και η  $\Phi^{-1}(p)$

$$\Phi: \mathbb{R} \rightarrow [0,1] \quad \Phi(x) = P(Z_i \leq x)$$

$$\Phi^{-1}: (0,1) \rightarrow \mathbb{R}$$



- ```
> plot(pnorm, xlim=c(-3, 3), ylim=c(-3, 3), col="blue", lwd=2)
> curve(qnorm(x), xlim=c(0.001, 0.999), add=T, col="red", lty=2, lwd=2)
> legend(-2, 2, legend=c("pnorm", "qnorm"), col=c("blue", "red"), lty=c(1, 2), lwd=2)
```

# Η α.σ.κ. της $N(0,1)$ , $\Phi(x)$ , η $\Phi^{-1}(p)$ και τα $z_\alpha$

$$\Phi(z_\alpha) = 1 - \alpha$$

$$z_\alpha = \Phi^{-1}(1 - \alpha)$$

$$z_\alpha = \text{qnorm}(1 - \alpha)$$

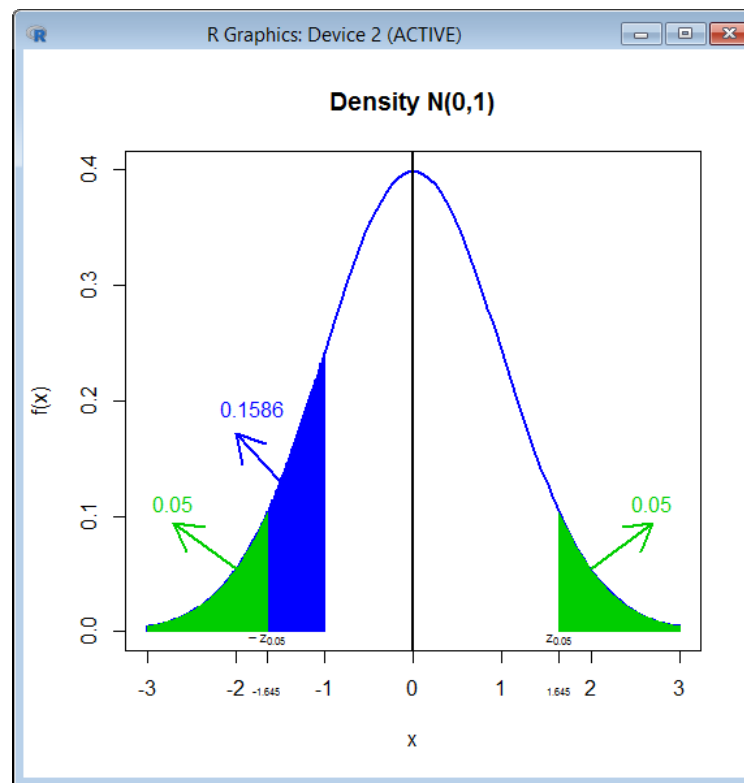
$$\Phi(-z_\alpha) = 1 - \Phi(z_\alpha) = \alpha$$

$$-z_\alpha = \Phi^{-1}(\alpha)$$

$$-z_\alpha = \text{qnorm}(\alpha)$$

$$z_\alpha = -\text{qnorm}(\alpha)$$

```
> pnorm(-2)
[1] 0.02275013
> pnorm(-1)
[1] 0.1586553
> pnorm(c(0, 1, 2, 3))
[1] 0.5000000 0.8413447 0.9772499 0.9986501
> qnorm(0.05)
[1] -1.644854
> qnorm(0.5) #median
[1] 0
> qnorm(0.95)
[1] 1.644854
```



$$\alpha = 0.1 \quad \delta.ε. \quad \sigma.ε \quad 90\%$$

$$z_{\alpha/2} = z_{0.05}$$

περίπου

$$\Phi(1.645) = 0.95$$

$$\Rightarrow z_{0.05} = 1.645$$

( $z_{0.25}$  είναι το  
τρίτο τεταρτημόριο της κατανομής)

$z_{0.75}$  είναι το πρώτο τεταρτημόριο } των  
 $z_{0.5}$  είναι η διάμεσος } κατανομών

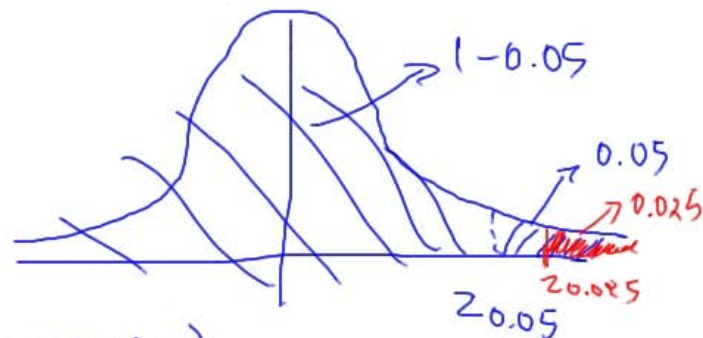
$$\alpha = 0.05 \quad \delta.ε. \quad \sigma.ε \quad 95\%$$

$$z_{\alpha/2} = z_{0.025}$$

$$\Phi(1.96) = 0.975$$

$$\Rightarrow z_{0.025} = 1.96$$

$$\Phi(z_{0.05}) = 1 - 0.05 = 0.95$$



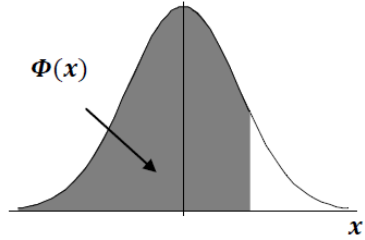
$$z_{0.05} < z_{0.025}$$

$$\Phi(z_{0.025}) = 1 - 0.025 = 0.975$$

| $x$ | .00   | .01   | .02   | .03   | .04   | .05   | .06   | .07   | .08   | .09   |
|-----|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5674 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9279 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .9871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .9893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |
| 3.0 | .9987 | .9987 | .9987 | .9988 | .9988 | .9989 | .9989 | .9989 | .9990 | .9990 |

Πίνακας τιμών της σ.κ.  $\Phi$  της τυπικής κανονικής κατανομής ( $x = 0, .01, .02, \dots, 3.49$ )

$$\Phi(x) = P(X \leq x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$



| $x$ | .00   | .01   | .02   | .03   | .04   | .05   | .06   | .07   | .08   | .09   |
|-----|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5674 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9279 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .9871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .9893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |
| 3.0 | .9987 | .9987 | .9987 | .9988 | .9988 | .9989 | .9989 | .9989 | .9990 | .9990 |

$$\Phi(z_{0.05}) = 0.95,$$

$$\Phi(1.645) = 0.95 \Rightarrow z_{0.05} = 1.645,$$

$$\Phi(z_{0.025}) = 0.975, \text{ όμως}$$

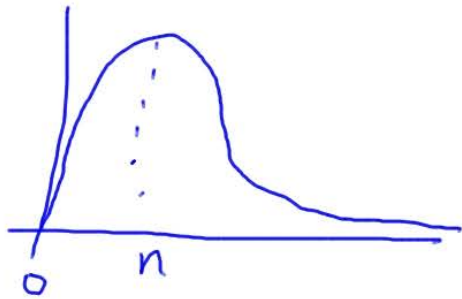
$$\Phi(1.96) = 0.975)$$

$$2) \cdot X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\sum_{i=1}^n \left( \frac{X_i - \mu}{\sigma} \right)^2 \sim \chi_n^2 \quad \chi^2 \text{ κατανομή με } n \text{ βαθμούς ελευθερίας}$$

$$X_i \sim N(\mu, \sigma^2) \quad Z_i = \frac{X_i - \mu}{\sigma} \sim N(0, 1) \quad Z_i^2 = \left( \frac{X_i - \mu}{\sigma} \right)^2 \sim \chi_1^2$$

$$\text{Εάν } W_i \sim \chi_1^2 \text{ και είναι ανεξάρτητες } W = \sum_{i=1}^n W_i \sim \chi_{1+\dots+1}^2 \equiv \chi_n^2$$



$$E(W) = n \quad \text{var}(W) = 2n$$

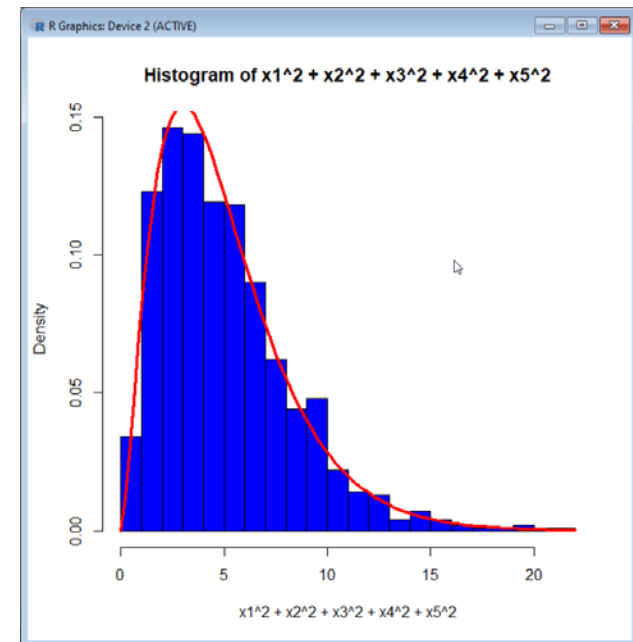
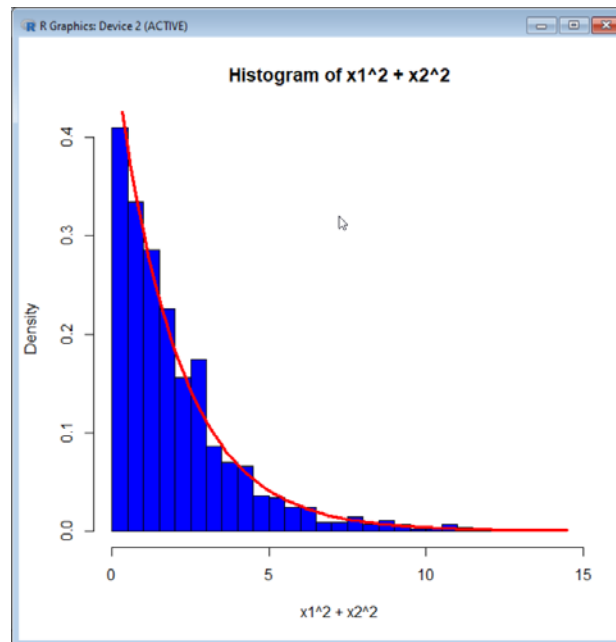
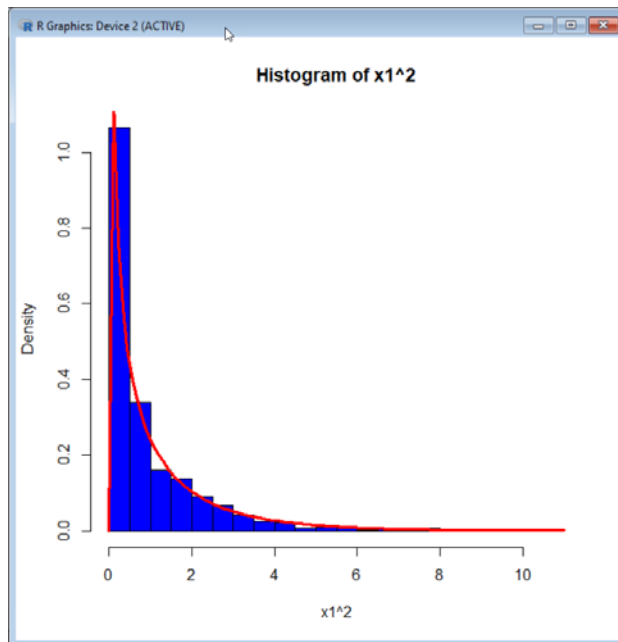
$$(n-1) \frac{\overline{S}^2}{\sigma^2} = \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{\sigma^2} \sim \chi_{n-1}^2 \quad **$$

$$\sum_{i=1}^n \frac{(X_i - \mu)^2}{\sigma^2} = \underbrace{\sum_{i=1}^n \frac{(X_i - \bar{X})^2}{\sigma^2}}_{\text{ανεξ}} + \left( \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \right)^2 \quad \chi_n^2 \equiv \chi_{n-1}^2 + \chi_1^2$$

```

> x1<-rnorm(1000)
> hist(x1^2,nclass=25,col=4,prob=TRUE)
> curve(dchisq(x,df=1),lwd=3,col="red",add=TRUE)
>
> x2<-rnorm(1000)
> hist(x1^2+x2^2,nclass=25,col=4,prob=TRUE)
> curve(dchisq(x,df=2),lwd=3,col="red",add=TRUE)
>
> x3<-rnorm(1000)
> x4<-rnorm(1000)
> x5<-rnorm(1000)
> hist(x1^2+x2^2+x3^2+x4^2+x5^2,nclass=25,col=4,prob=TRUE)
> curve(dchisq(x,df=5),lwd=3,col="red",add=TRUE)
>
> round(c(mean(x1^2),mean(x1^2+x2^2),mean(x1^2+x2^2+x3^2),mean(x1^2+x2^2+x3^2+x4^2+x5^2)),2)
[1] 0.96 2.03 3.05 5.06
> round(c(var(x1^2),var(x1^2+x2^2),var(x1^2+x2^2+x3^2),var(x1^2+x2^2+x3^2+x4^2+x5^2)),2)
[1] 1.85 4.06 6.40 10.41

```



$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

---

$$E\left(\frac{(n-1)S^2}{\sigma^2}\right) = n-1$$

$$\text{var}\left(\frac{(n-1)S^2}{\sigma^2}\right) = 2(n-1)$$

$$\frac{(n-1)}{\sigma^2} E(S^2) = n-1$$

$$\frac{(n-1)^2}{(\sigma^2)^2} \text{var}(S^2) = 2(n-1)$$

$$E(S^2) = \sigma^2$$

$$\text{var}(S^2) = \frac{2\sigma^4}{n-1}$$

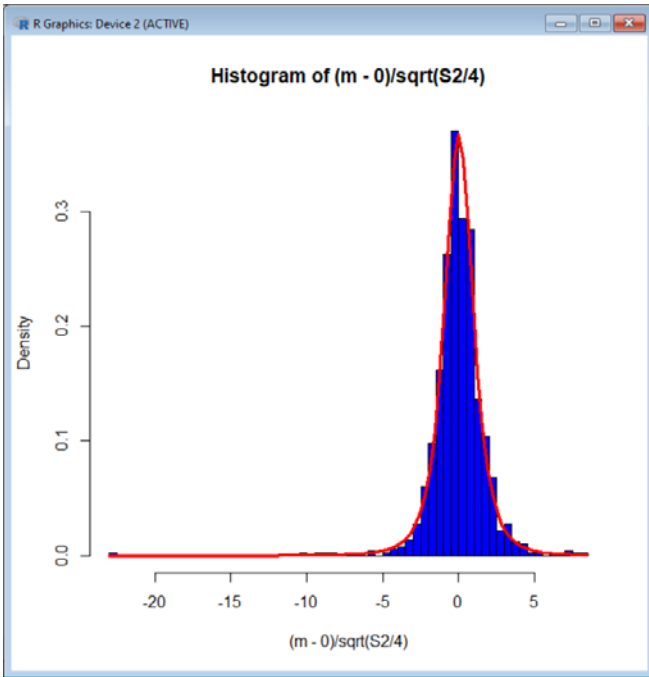
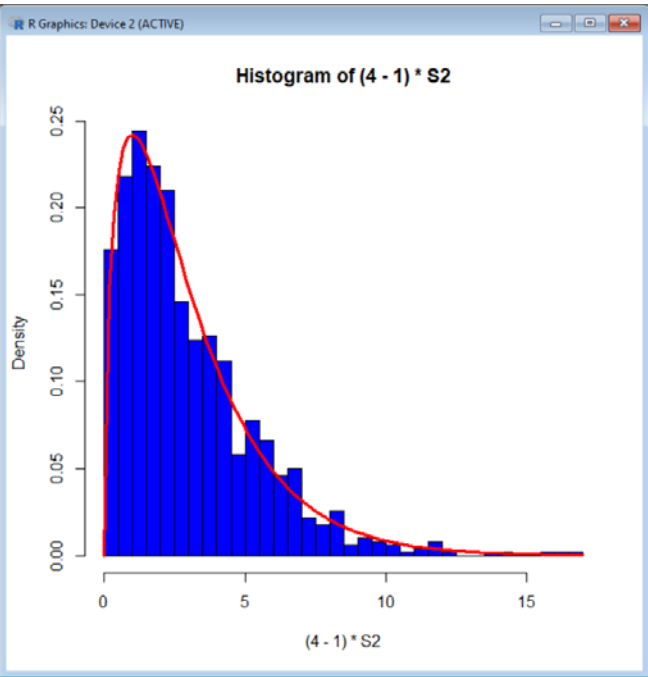
γνωστό  
ισχύει για οποιαδήποτε  
κατανομή

αφορά παρατηρήσεις  
από κανονική κατανομή

```

> X1<-rnorm(5)
> X2<-rnorm(5)
> X3<-rnorm(5)
>
> rbind(X1,X2,X3)
      [,1]      [,2]      [,3]      [,4]      [,5]
X1  0.1303082  1.0434044  0.42225412 -1.5889194  0.6035728
X2  0.8038980  0.1791883  0.04841418 -0.5678942  0.2770057
X3 -0.2032669 -2.0538466  1.36738044  0.4260789  0.6835301
>
> m<-apply(rbind(X1,X2,X3),2,mean)
> m
[1]  0.2436464 -0.2770846  0.6126829 -0.5769116  0.5213695
> S2<-apply(rbind(X1,X2,X3),2,var)
> S2
[1]  0.26322944  2.55437979  0.46211535  1.01511544  0.04638354

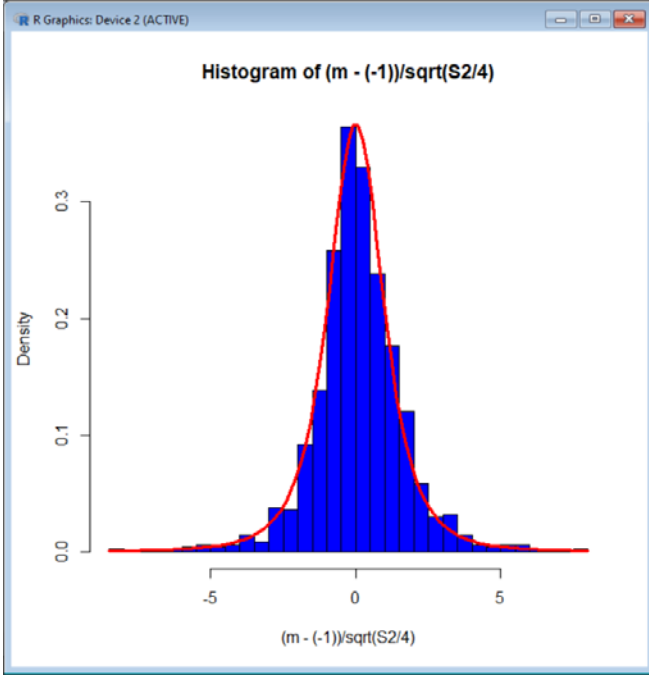
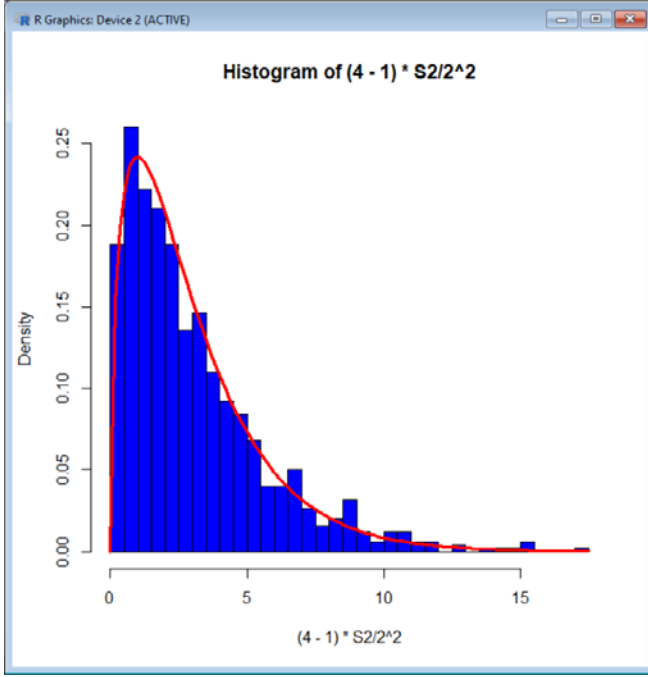
```



```

> X1<-rnorm(1000)
> X2<-rnorm(1000)
> X3<-rnorm(1000)
> X4<-rnorm(1000)
>
> m<-apply(rbind(X1,X2,X3,X4),2,mean)
> S2<-apply(rbind(X1,X2,X3,X4),2,var)
>
> hist((4-1)*S2/1^2,nclass=50,col=4,prob=TRUE)
> curve(dchisq(x,df=3),lwd=3,col="red",add=TRUE)
>
> hist((m-0)/sqrt(S2/4),nclass=50,col=4,prob=TRUE)
> curve(dt(x,df=3),lwd=3,col="red",add=TRUE)

```



```

> X1<-rnorm(1000,mean=-1,sd=2)
> X2<-rnorm(1000,mean=-1,sd=2)
> X3<-rnorm(1000,mean=-1,sd=2)
> X4<-rnorm(1000,mean=-1,sd=2)
>
> hist((4-1)*S2/2^2,nclass=50,col=4,prob=TRUE)
> hist((m-(-1))/sqrt(S2/4),nclass=50,col=4,prob=TRUE)

```



# Η α.σ.κ. της $\chi^2$ , $F(x)$ , η $F^{-1}(p)$ και τα $\chi^2_{v,\alpha}$

$X_1, X_2, \dots, X_n$  τυχαίο δείγμα από  $N(\mu, \sigma^2)$

$$\sum_{i=1}^n \frac{(X_i - \mu)^2}{\sigma^2} \sim \chi^2_n$$

$$(n-1) \frac{S^2}{\sigma^2} = \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{\sigma^2} \sim \chi^2_{n-1}$$

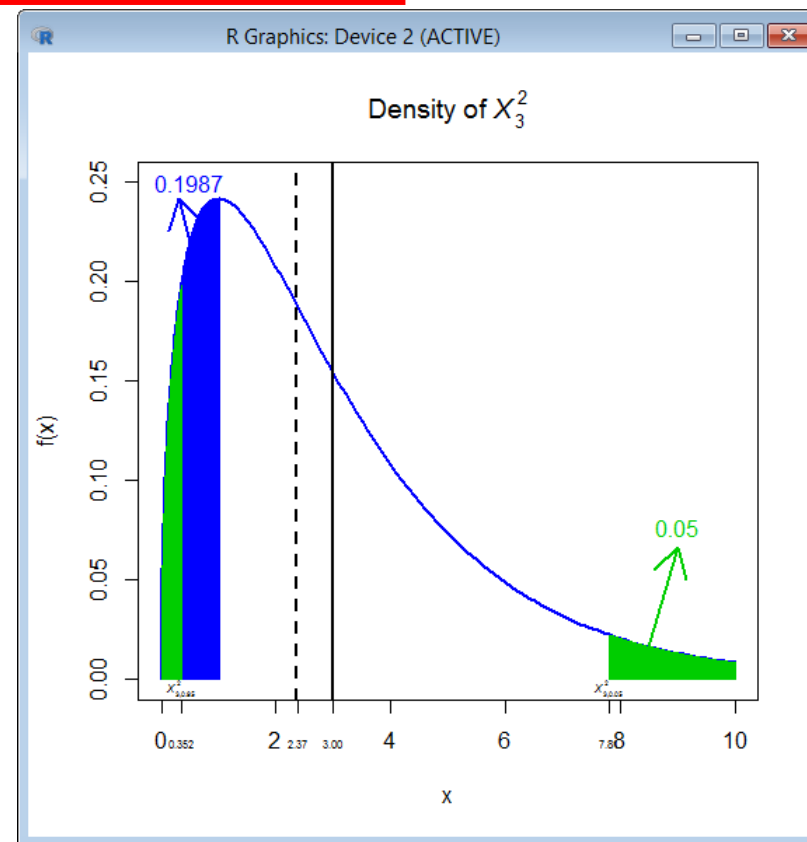
$$F(\chi^2_{v,\alpha}) = 1 - \alpha$$

$$\chi^2_{v,\alpha} = F^{-1}(1 - \alpha) \quad \chi^2_{v,\alpha} = \text{qchisq}(1 - \alpha, \text{df} = v)$$

$$F(\chi^2_{v,1-\alpha}) = \alpha$$

$$\chi^2_{v,1-\alpha} = F^{-1}(\alpha) \quad \chi^2_{v,1-\alpha} = \text{qchisq}(\alpha, v)$$

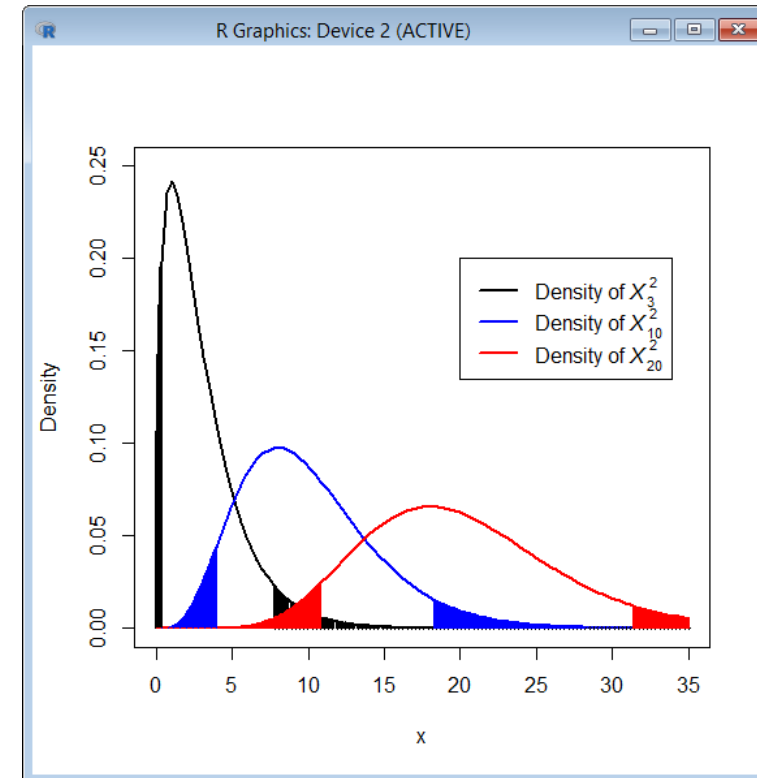
```
> v<-3
> pchisq(-1, df=v)
[1] 0
> pchisq(c(1, 3, 9), v)
[1] 0.1987480 0.6083748 0.9707091
> qchisq(0.05, v)
[1] 0.3518463
> qchisq(0.5, v) #median
[1] 2.365974
> qchisq(0.95, v)
[1] 7.814728
```



# Τα ποσοστιαία σημεία $\chi^2_{v,\alpha}$ για διάφορα $v$

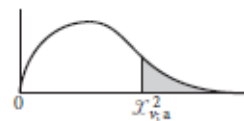
```
> #df=10
> v<-10
> pchisq(c(1,10,23),v)
[1] 0.0001721156 0.5595067149 0.9892534216
> qchisq(0.05,v)
[1] 3.940299
> qchisq(0.5,v) #median
[1] 9.341818
> qchisq(0.95,v)
[1] 18.30704

> #df=20
> v<-20
> qchisq(0.05,v)
[1] 10.85081
> pchisq(c(1,20,36),v)
[1] 1.709670e-10 5.420703e-01 9.846189e-01
> qchisq(0.5,v) #median
[1] 19.33743
> qchisq(0.95,v)
[1] 31.41043
```



Πίνακας: Τιμών  $\chi^2_{v; a}$  της  $\chi^2$  κατανομής για τις οποίες

$$P(\chi^2 > \chi^2_{v; a}) = a$$



| $\beta.ε.$ | $a=0,995$ | $a=0,990$ | $a=0,975$ | $a=0,950$ | $a=0,900$ |
|------------|-----------|-----------|-----------|-----------|-----------|
| 1          | 0,0000393 | 0,0001571 | 0,0009821 | 0,0039321 | 0,0157908 |
| 2          | 0,0100251 | 0,0201007 | 0,0506356 | 0,102587  | 0,210720  |
| 3          | 0,0717212 | 0,114832  | 0,215795  | 0,351846  | 0,584375  |
| 4          | 0,206990  | 0,297110  | 0,484419  | 0,710721  | 1,063623  |
| 5          | 0,411740  | 0,554300  | 0,831211  | 1,145476  | 1,61031   |
| 6          | 0,675727  | 0,872085  | 1,237347  | 1,63539   | 2,20413   |
| 7          | 0,989265  | 1,239043  | 1,68987   | 2,16735   | 2,83311   |
| 8          | 1,344419  | 1,646482  | 2,17973   | 2,73264   | 3,48954   |
| 9          | 1,734926  | 2,087912  | 2,70039   | 3,32511   | 4,16816   |
| 10         | 2,15585   | 2,55821   | 3,24697   | 3,94030   | 4,86518   |
| 11         | 2,60321   | 3,05347   | 3,81575   | 4,57481   | 5,57779   |
| 12         | 3,07382   | 3,57056   | 4,40379   | 5,22603   | 6,30380   |
| 13         | 3,56503   | 4,10691   | 5,00874   | 5,89186   | 7,04150   |
| 14         | 4,07468   | 4,66043   | 5,62872   | 6,57063   | 7,78953   |
| 15         | 4,60094   | 5,22935   | 6,26214   | 7,26094   | 8,54675   |
| 16         | 5,14224   | 5,81221   | 6,90766   | 7,96164   | 9,31223   |
| 17         | 5,69724   | 6,40776   | 7,56418   | 8,67176   | 10,0852   |
| 18         | 6,26481   | 7,01491   | 8,23075   | 9,39046   | 10,8649   |
| 19         | 6,84398   | 7,63273   | 8,90655   | 10,1170   | 11,6509   |
| 20         | 7,43386   | 8,26040   | 9,59083   | 10,8508   | 12,4426   |
| 21         | 8,03366   | 8,89720   | 10,28293  | 11,5913   | 13,2396   |
| 22         | 8,64272   | 9,54249   | 10,9823   | 12,3380   | 14,0415   |
| 23         | 9,26042   | 10,19567  | 11,6885   | 13,0905   | 14,8479   |
| 24         | 9,88623   | 10,8564   | 12,4011   | 13,8484   | 15,6587   |
| 25         | 10,5197   | 11,5240   | 13,1197   | 14,6114   | 16,4734   |
| 26         | 11,1603   | 12,1981   | 13,8439   | 15,3791   | 17,2919   |
| 27         | 11,8076   | 12,8786   | 14,5733   | 16,1513   | 18,1138   |
| 28         | 12,4613   | 13,5648   | 15,3079   | 16,9279   | 18,9392   |
| 29         | 13,1211   | 14,2565   | 16,0471   | 17,7083   | 19,7677   |
| 30         | 13,7867   | 14,9535   | 16,7908   | 18,4926   | 20,5992   |
| 40         | 20,7065   | 22,1643   | 24,4331   | 26,5093   | 29,0505   |
| 50         | 27,9907   | 29,7067   | 32,3574   | 34,7642   | 37,6886   |
| 60         | 35,5346   | 37,4848   | 40,4817   | 43,1879   | 46,4589   |
| 70         | 43,2752   | 45,4418   | 48,7576   | 51,7393   | 55,3290   |
| 80         | 51,1720   | 53,5400   | 57,1532   | 60,3915   | 64,2778   |
| 90         | 59,1963   | 61,7541   | 65,6466   | 69,1260   | 73,2912   |
| 100        | 67,3276   | 70,0648   | 74,2219   | 77,9295   | 82,3581   |

| $a=0,10$ | $a=0,05$ | $a=0,025$ | $a=0,010$ | $a=0,005$ | $\beta.ε.$ |
|----------|----------|-----------|-----------|-----------|------------|
| 2,70554  | 3,84146  | 5,02389   | 6,63490   | 7,87944   | 1          |
| 4,60517  | 5,99147  | 7,37776   | 9,21034   | 10,5966   | 2          |
| 6,25139  | 7,81473  | 9,34840   | 11,3449   | 12,8381   | 3          |
| 7,77944  | 9,48773  | 11,1433   | 13,2767   | 14,8602   | 4          |
| 9,23635  | 11,0705  | 12,8325   | 15,0863   | 16,7496   | 5          |
| 10,6446  | 12,5916  | 14,4494   | 16,8119   | 18,5476   | 6          |
| 12,0170  | 14,0671  | 16,0128   | 18,4753   | 20,2777   | 7          |
| 13,3616  | 15,5073  | 17,5346   | 20,0902   | 21,9550   | 8          |
| 14,6837  | 16,9190  | 19,0228   | 21,6660   | 23,5893   | 9          |
| 15,9871  | 18,3070  | 20,4831   | 23,2093   | 25,1882   | 10         |
| 17,2750  | 19,6751  | 21,9200   | 24,7250   | 26,7569   | 11         |
| 18,5494  | 21,0261  | 23,3367   | 26,2170   | 28,2995   | 12         |
| 19,8119  | 22,3621  | 24,7356   | 27,6883   | 29,8194   | 13         |
| 21,0642  | 23,6848  | 26,1190   | 29,1413   | 31,3193   | 14         |
| 22,3072  | 24,9958  | 27,4884   | 30,5779   | 32,8013   | 15         |
| 23,5418  | 26,2962  | 28,8454   | 31,9999   | 34,2672   | 16         |
| 24,7690  | 27,5871  | 30,1910   | 33,4087   | 35,7185   | 17         |
| 25,9894  | 28,8693  | 31,5264   | 34,8053   | 37,1564   | 18         |
| 27,2036  | 30,1435  | 32,8523   | 36,1908   | 38,5822   | 19         |
| 28,4120  | 31,4104  | 34,1696   | 37,5662   | 39,9968   | 20         |
| 29,6151  | 32,6705  | 35,4789   | 38,9321   | 41,4010   | 21         |
| 30,8133  | 33,9244  | 36,7807   | 40,2894   | 42,7956   | 22         |
| 32,0069  | 35,1725  | 38,0757   | 41,6384   | 44,1813   | 23         |
| 33,1963  | 36,4151  | 39,3641   | 42,9798   | 45,5585   | 24         |
| 34,3816  | 37,6525  | 40,6465   | 44,3141   | 46,9278   | 25         |
| 35,5631  | 38,8852  | 41,9232   | 45,6417   | 48,2899   | 26         |
| 36,7412  | 40,1133  | 43,1944   | 46,9630   | 49,6449   | 27         |
| 37,9159  | 41,3372  | 44,4607   | 48,2782   | 50,9933   | 28         |
| 39,0875  | 42,5569  | 45,7222   | 49,5879   | 52,3356   | 29         |
| 40,2560  | 43,7729  | 46,9792   | 50,8922   | 53,6720   | 30         |
| 51,8050  | 55,7585  | 59,3417   | 63,6907   | 66,7659   | 40         |
| 63,1671  | 67,5048  | 71,4202   | 76,1539   | 79,4900   | 50         |
| 74,3970  | 79,0819  | 83,2976   | 88,3794   | 91,9517   | 60         |
| 85,5271  | 90,5312  | 95,0231   | 100,425   | 104,215   | 70         |
| 96,5782  | 101,879  | 106,629   | 112,329   | 116,321   | 80         |
| 107,565  | 113,145  | 118,136   | 124,116   | 128,299   | 90         |
| 118,498  | 124,342  | 129,561   | 135,807   | 140,169   | 100        |

Για  $v > 100$ ,  $\chi^2_{v; a} = \frac{1}{2} (z_a + \sqrt{2v-1})^2$

$\sigma^2$  άγνωστο

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1)$$

$$(n-1) \frac{S^2}{\sigma^2} \sim \chi_{n-1}^2$$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

$$(X_i \sim N(\mu, \sigma^2) \text{ ανεξ})$$

$$\frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}} = \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}}{\sqrt{\frac{(n-1)S^2}{\sigma^2} / n-1}} \sim \frac{N(0, 1)}{\sqrt{\frac{\chi_{n-1}^2}{n-1}}} \quad \text{ανεξ}$$

$$\equiv t_{n-1} \quad (\text{νέα κατανομή})$$

συμμετρική ως προς το 0  
και μοιάζει με την  $N(0, 1)$

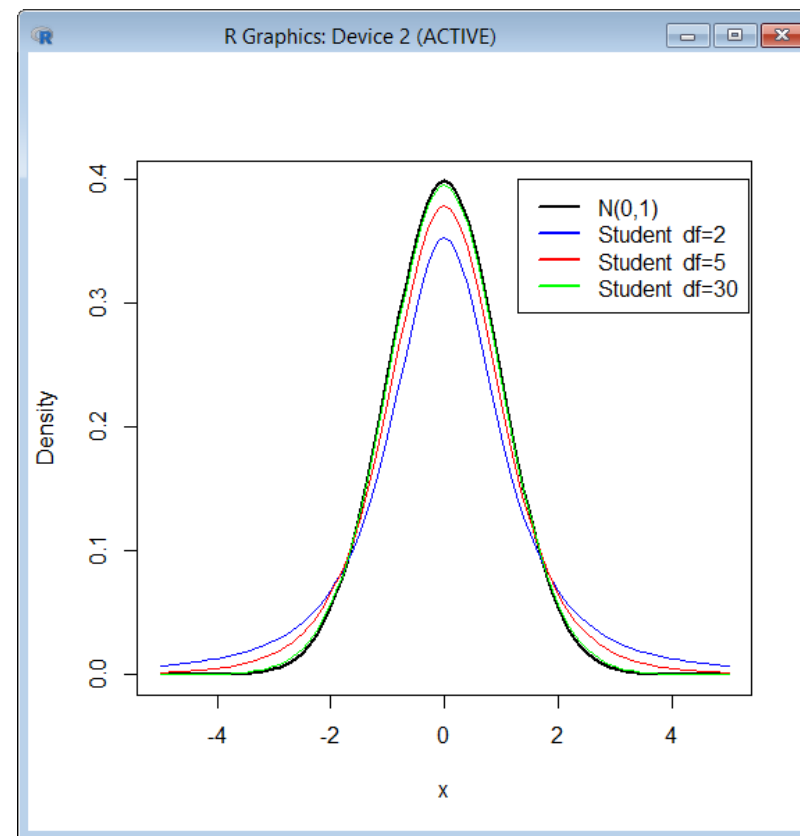
για μικρό  $n$  είναι πιο  
"πλάτια", για  $n \gg \gg 0$   $t_n \approx N(0, 1)$

# Τυπική κανονική κατανομή και Student's t κατανομή

$X_1, X_2, \dots, X_n$  τυχαίο δείγμα από  $N(\mu, \sigma^2)$

$$\frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}} \sim t_{n-1}$$

$$\frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}} = \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}}{\sqrt{\frac{S^2}{\sigma^2}}} = \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}}{\sqrt{\frac{(n-1)S^2}{(n-1)\sigma^2}}} \equiv \frac{N(0,1)}{\sqrt{\frac{\chi^2_{n-1}}{n-1}}} \Rightarrow \bar{X} \text{ και } S^2 \text{ ανεξάρτητες}$$



```
> plot(dnorm, xlim=c(-5,5), lwd=2, ylab="Density")
> curve(dt(x, df=2), xlim=c(-5,5), col="blue", add=T)
> curve(dt(x, df=5), xlim=c(-5,5), col="red", add=T)
> curve(dt(x, df=30), xlim=c(-5,5), col="green", add=T)
> legend(1.3, 0.4, legend=c("N(0,1)", "Student df=2",
+ "Student df=5", "Student df=30"), col=c(1, "blue", "red", "green"),
```

# Τα ποσοστιαία σημεία $t_{v;\alpha}$ και βαθμοί ελευθερίας

$$F(t_{v,\alpha})=1-\alpha$$

$$t_{v,\alpha}=F^{-1}(1-\alpha)$$

$$F(-t_{v,\alpha})=\alpha$$

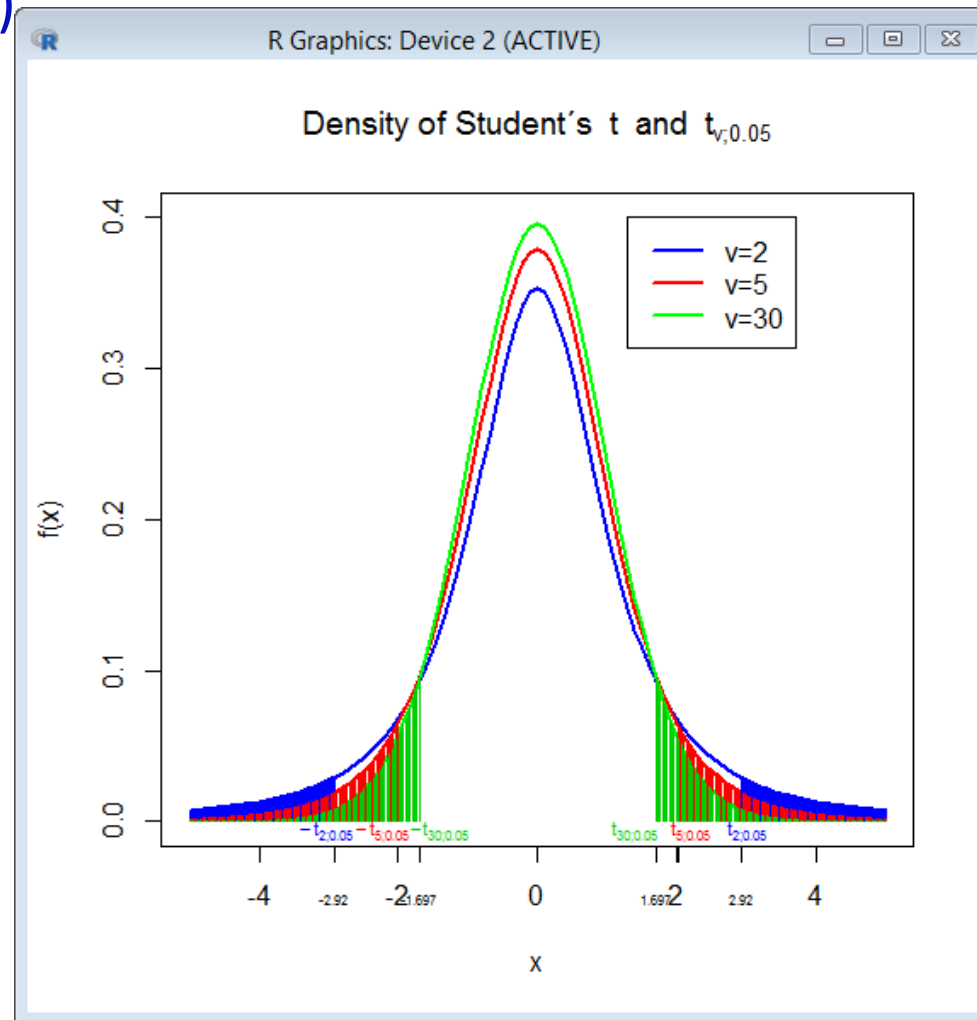
$$-t_{v,\alpha}=F^{-1}(\alpha)$$

$$t_{v,\alpha}=qt(1-\alpha, df=v)$$

$$-t_{v,\alpha}=qt(\alpha, v)$$

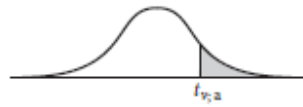
$$t_{v,\alpha}=-qt(\alpha, v)$$

```
> pt(-2, df=2)
[1] 0.09175171
> qt(0.05, df=2)
[1] -2.919986
> qt(0.05, df=5)
[1] -2.015048
> qt(0.05, df=30)
[1] -1.697261
> qt(0.95, df=c(2, 5, 30, 50, 100))
[1] 2.919986 2.015048 1.697261
[4] 1.675905 1.660234
> qnorm(0.95)
[1] 1.644854
```



Πίνακας: Τιμών  $t_{v;a}$  της  $t_v$ -κατανομής ώστε

$$P(t_v > t_{v;a}) = a$$



| $\beta.ε.$ | $a=0,10$ | $a=0,05$ | $a=0,025$ | $a=0,010$ | $a=0,005$ |
|------------|----------|----------|-----------|-----------|-----------|
| 1          | 3,078    | 6,314    | 12,706    | 31,821    | 63,657    |
| 2          | 1,886    | 2,920    | 4,303     | 6,965     | 9,925     |
| 3          | 1,638    | 2,353    | 3,182     | 4,541     | 5,841     |
| 4          | 1,533    | 2,132    | 2,776     | 3,747     | 4,604     |
| 5          | 1,476    | 2,015    | 2,571     | 3,365     | 4,032     |
| 6          | 1,440    | 1,943    | 2,447     | 3,143     | 3,707     |
| 7          | 1,415    | 1,895    | 2,365     | 2,998     | 3,499     |
| 8          | 1,397    | 1,860    | 2,306     | 2,896     | 3,355     |
| 9          | 1,383    | 1,833    | 2,262     | 2,821     | 3,250     |
| 10         | 1,372    | 1,812    | 2,228     | 2,764     | 3,169     |
| 11         | 1,363    | 1,796    | 2,201     | 2,718     | 3,106     |
| 12         | 1,356    | 1,782    | 2,179     | 2,681     | 3,055     |
| 13         | 1,350    | 1,771    | 2,160     | 2,650     | 3,012     |
| 14         | 1,345    | 1,761    | 2,145     | 2,624     | 2,977     |
| 15         | 1,341    | 1,753    | 2,131     | 2,602     | 2,947     |
| 16         | 1,337    | 1,746    | 2,120     | 2,583     | 2,921     |
| 17         | 1,333    | 1,740    | 2,110     | 2,567     | 2,898     |
| 18         | 1,330    | 1,734    | 2,101     | 2,552     | 2,878     |
| 19         | 1,328    | 1,729    | 2,093     | 2,539     | 2,861     |
| 20         | 1,325    | 1,725    | 2,086     | 2,528     | 2,845     |
| 21         | 1,323    | 1,721    | 2,080     | 2,518     | 2,831     |
| 22         | 1,321    | 1,717    | 2,074     | 2,508     | 2,819     |
| 23         | 1,319    | 1,714    | 2,069     | 2,500     | 2,807     |
| 24         | 1,318    | 1,711    | 2,064     | 2,492     | 2,797     |
| 25         | 1,316    | 1,708    | 2,060     | 2,485     | 2,787     |
| 26         | 1,315    | 1,706    | 2,056     | 2,479     | 2,779     |
| 27         | 1,314    | 1,703    | 2,052     | 2,473     | 2,771     |
| 28         | 1,313    | 1,701    | 2,048     | 2,467     | 2,763     |
| 29         | 1,311    | 1,699    | 2,045     | 2,462     | 2,756     |
| $\infty$   | 1,282    | 1,645    | 1,960     | 2,326     | 2,576     |

$$F(t_{v,0.05}) = 1 - 0.05 = 0.95$$

$$t_{v,0.05} = F^{-1}(0.95)$$

$$F(-t_{v,0.05}) = 0.05$$

$$-t_{v,0.05} = F^{-1}(0.05)$$

$$t_{2,0.05} = 2.920$$

$$t_{5,0.05} = 2.015$$

$$t_{9,0.05} = ( \text{ή } t_9(0.05) ) 1.833$$

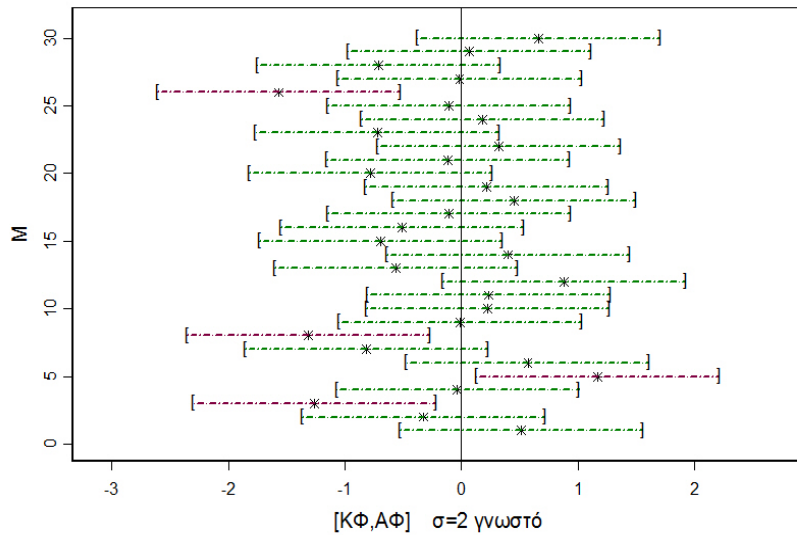
$$t_{30,0.05} = 1.697 \approx z_{0.05} = 1.645$$

# Διαστήματα Εμπιστοσύνης

σ.ε. 90%(=1-α, α=0.1, α/2=0.05) για την μέση τιμή με γνωστή (σ=2) και άγνωστη διασπορά

Χρησιμοποιήθηκαν 30 δείγματα, μεγέθους n=10 έκαστον, από κανονική κατανομή N(μ=0,σ²=2²=4)

και τα ποσοστιαία σημεία: από την τυπική κανονική  $z_{0.05}=1.645$  και από την Student (με n-1=9 β.ε.)  $t_{9,0.05}=t_9(0.05)=1.833$



$$\left[ \bar{X} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \right]$$

$$\left[ \bar{X} - t_{n-1}(\alpha/2) \frac{S}{\sqrt{n}}, \bar{X} + t_{n-1}(\alpha/2) \frac{S}{\sqrt{n}} \right]$$

